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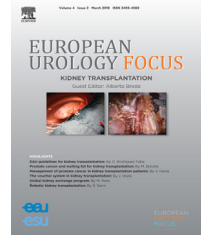
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Urothelial Cancer

A Preoperative Nomogram to Predict Renal Function Insufficiency for Cisplatin-based Adjuvant Chemotherapy Following Minimally Invasive Radical Nephroureterectomy (ROBUUST Collaborative Group)

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Abstract

Background: Postoperative renal function impairment represents a main limitation for delivering adjuvant chemotherapy after radical nephroureterectomy (RNU).

Objective: To create a model predicting renal function decline after minimally invasive RNU.

Design, setting, and participants: A total of 490 patients with nonmetastatic UTUC who underwent minimally invasive RNU were identified from a collaborative database including 17 institutions worldwide (February 2006 to March 2020). Renal function insufficiency for cisplatin-based regimen was defined as estimated glomerular filtration rate (eGFR) <50 ml/min/1.73 m² at 3 mo after RNU. Patients with baseline eGFR >50 ml/min/1.73 m² (n = 361) were geographically divided into a training set (n = 226) and an independent external validation set (n = 135) for further analysis.

Outcome measurements and statistical analysis: Using transparent reporting of a multivariable prediction model for individual prognosis or diagnosis (TRIPOD) guidelines, a nomogram to predict postoperative eGFR <50 ml/min/1.73 m² was built based on

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the coefficients of the least absolute shrinkage and selection operation (LASSO) logistic regression. The discrimination, calibration, and clinical use of the nomogram were investigated.

Results and limitations: The model that incorporated age, body mass index, preoperative eGFR, and hydronephrosis was developed with an area under the curve of 0.771, which was confirmed to be 0.773 in the external validation set. The calibration curve demonstrated good agreement. Besides, the model was converted into a risk score with a cutoff value of 0.583, and the difference between the low- and high-risk groups both in overall death risk (hazard ratio [HR]: 4.59, $p < 0.001$) and cancer-specific death risk (HR: 5.19, $p < 0.001$) was statistically significant. The limitation mainly lies in its retrospective design.

Conclusions: A nomogram incorporating immediately available clinical variables can accurately predict renal insufficiency for cisplatin-based adjuvant chemotherapy after minimally invasive RNU and may serve as a tool facilitating patient selection.

Patient summary: We have developed a model for the prediction of renal function loss after radical nephroureterectomy to facilitate patient selection for perioperative chemotherapy.

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1. Introduction

The value of adjuvant platinum-based chemotherapy for patients undergoing radical nephroureterectomy (RNU) with curative intent has recently been described with level I evidence, while a subgroup analysis showed that cisplatin-treated patients benefited significantly but carboplatin-treated patients did not [1,2]. Moreover, cisplatin use is found to be an independent favorable prognostic factor, and this benefit is independent of baseline characteristics or comorbidities [3].

The main limitation of using adjuvant chemotherapy remains the limited ability to deliver cisplatin-based regimen following RNU, given that the surgical procedure is likely to impact renal function, which is the most common reason of cisplatin-based treatment ineligibility [3–5]. Therefore, tools predicting renal functional decline would facilitate perioperative management planning and help in the selection of those patients who could benefit the most from neoadjuvant chemotherapy. Although there are reported clinical factors associated with worse renal functional outcomes [6,7], to the best of our knowledge, there are no existing validated nomograms for predicting renal function insufficiency for cisplatin-based adjuvant chemotherapy after RNU.

The aim of the present study was to develop and validate a model for the prediction of significant renal function reduction in a global multi-institutional dataset (ROBotic surgery for Upper tract Urothelial cancer Study—ROBUUST project), thereby allowing the identification of patients likely to be ineligible for cisplatin-based adjuvant chemotherapy after minimally invasive radical nephroureterectomy, and investigate its significance in prognostic risk stratification.

2. Patients and methods

2.1. Study population and variables

ROBUUST is a multinational, multicenter project including 17 institutions worldwide. A dataset of patients who underwent single-stage minimally invasive (robotic or laparoscopic) RNU for upper tract urothelial carcinoma (UTUC) was created. Institutional review board approval or exempt was obtained at each center. The study design for this project was not

unblinded until the completion of data collection. The purpose-built ROBUUST database including cases of minimally invasive RNU performed at participating centers from February 2006 to March 2020 was queried with the following exclusion criteria: (1) preoperative metastatic diseases, (2) anatomically single kidney, and (3) missing data in predictors.

The following variables of interest were included for analysis:

- 1 Demographics, smoking history, American Society of Anesthesiologists (ASA) score, hypertension, diabetes, presence of hematuria, history of bladder cancer, preoperative serum creatinine (SCr) recorded most recent to the surgery of RNU (estimated glomerular filtration rate [eGFR] calculated by the Chronic Kidney Disease Epidemiology Collaboration formula [8]), tumor characteristics, and perioperative data
- 2 Functional data, including SCr and eGFR at postoperative day 1, discharge, postoperative 3 and 12 mo, and last follow-up
- 3 Oncological data, including tumor recurrence, tumor metastasis, cancer-specific mortality, and overall mortality

2.2. Outcome definition

Regarding the timing of postoperative eGFR for analysis, closest to 3 mo after surgery was selected to best approximate the measured SCr that would affect the delivery of adjuvant chemotherapy [1,5]. Although an eGFR of <60 ml/min/1.73 m² forms part of the Galsky definition of being cisplatin unfit [9], here we use the most recent criteria of eGFR <50 ml/min/1.73 m² according to UTUC-specific perioperative trials (POUT trial, etc.) [1,10].

2.3. Model development, validation, and clinical use analysis

According to the tutorials about clinical prediction models, nonrandom splitting of training and validation sets (eg, by centers) is considered more preferable, as it reduces the similarity of the two sets of patients and, generally, more individuals should be allocated to the training set [11,12]. Besides, geographic external validation is often possible with collaborative studies and is more meaningful than a standard cross-validation [13]. Thus, the sample was geographically separated into a training cohort (data from centers in the USA) and an external validation cohort (data from centers in China, Belgium, Italy, and Korea). The methodology of sample size estimation is provided in the Supplementary material. The least absolute shrinkage and selection operation (LASSO) regression model was used for the selection of variables in

the training set. Model development was performed according to the transparent reporting of a multivariable prediction model for individual prognosis or diagnosis (TRIPOD) guideline [14]. Additional methodology is provided in the Supplementary material.

2.4. Statistical analysis

The LASSO regression model was used with penalty parameter tuning that was conducted by ten-fold cross-validation based on 1 standard error criteria. A restricted cubic spline (RCS) analysis was adopted to assess the nonlinear relationship between the candidate variables and the predicted outcome. The decision curve analysis (DCA) method was used to evaluate the clinical utility of the presented nomogram. Overall death risk and cancer-specific death risk were obtained using the Kaplan-Meier method, and hazard ratios (HRs) with 95% confidence intervals (CIs) were calculated using a univariate Cox regression analysis. Statistical analyses were performed using R 3.6.0 software (<http://cran.r-project.org>). All tests were two sided, with a significance level set at $p < 0.05$.

3. Results

3.1. Baseline patient characteristics

A total of 490 patients were identified; out of them, 129 (26.3%) patients had preoperative eGFR < 50 ml/min/1.73 m², who were considered initially ineligible for cisplatin-based neoadjuvant or adjuvant chemotherapy. Finally, 361 patients were included for further analysis, with 226 cases in the training set and 135 cases in the validation set (Fig. 1).

Descriptive characteristics for this cohort, the training set (X), and the validation set (Y) are reported in Table 1. The median (interquartile range [IQR]) age at surgery of the whole, training, and validation sets was 70 (62, 77), 72 (63, 78), and 67 (60, 74.5) yr, respectively; 240 (66.5%), 141 (62.4%), and 99 (73.3%) patients were male in the respective sets. Ipsilateral hydronephrosis was found in 162

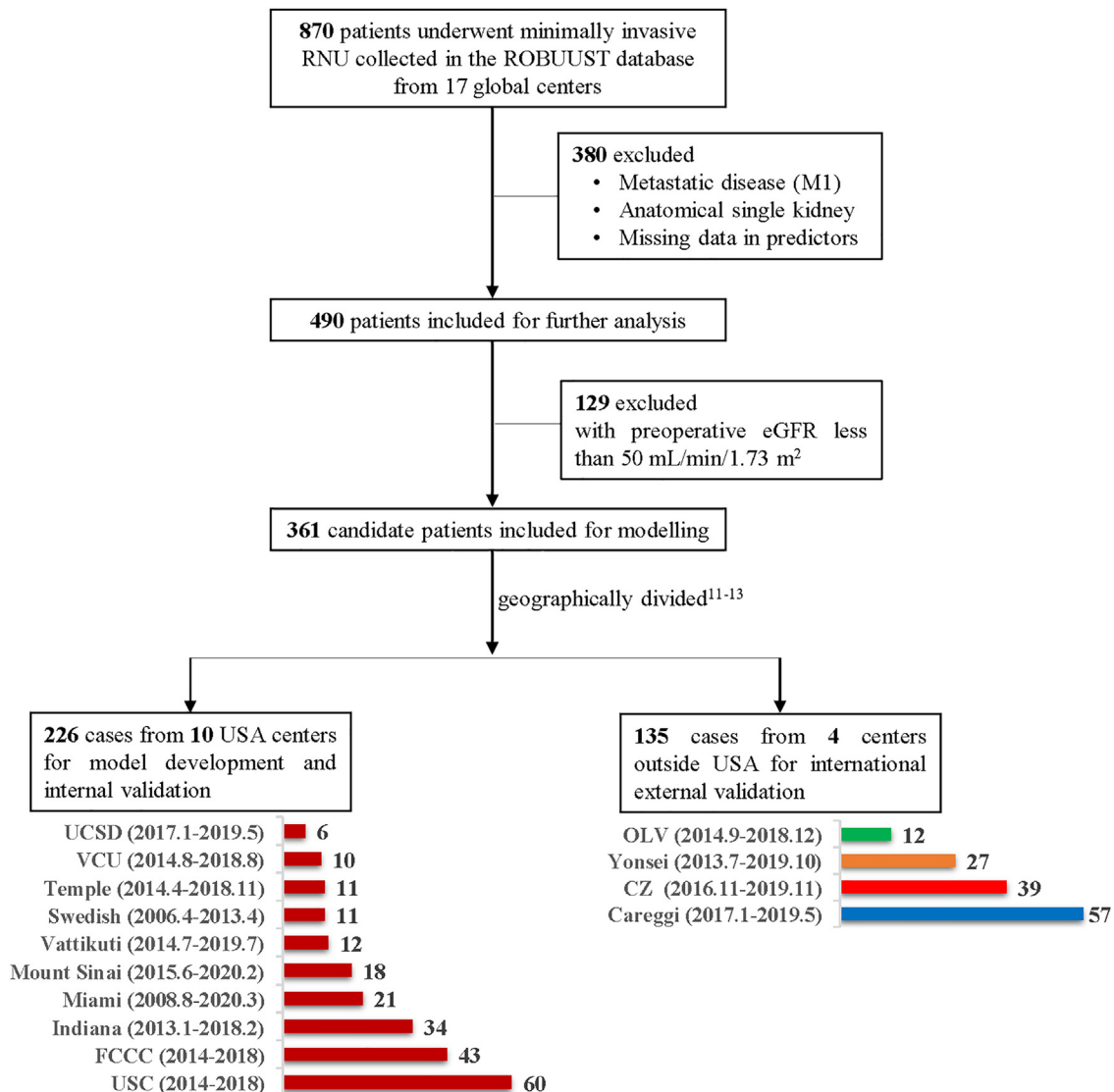


Fig. 1 – Diagram of study cohort selection steps. eGFR = estimated glomerular filtration rate; RNU = radical nephroureterectomy; ROBUUST = ROBOtic surgery for Upper tract Urothelial cancer Study.

Table 1 – Descriptive characteristics of the overall cohort, training set, and validation set

Variables	Overall (n = 361)	Training set (n = 226)	Validation set (n = 135)	p value ^a
Number of centers	14	10	4	
Age (yr)	70 (62, 77)	72 (63, 78)	67 (60, 74.5)	0.001
Race, n (%)				<0.001
Caucasian	237 (65.7)	168 (74.3)	69 (51.1)	
Black/Hispanic/Asian	117 (32.4)	51 (22.6)	66 (48.9)	
Other	7 (1.9)	7 (3.1)	0 (0)	
Gender, n (%)				0.033
Male	240 (66.5)	141 (62.4)	99 (73.3)	
Female	121 (33.5)	85 (37.6)	36 (26.7)	
BMI (kg/m ²)	26.2 (24.2, 29)	26.7 (24.4, 30.5)	25.7 (24, 27)	<0.001
Smoking history, n (%)	196 (54.3)	143 (63.3)	53 (39.3)	<0.001
ASA score, n (%)				<0.001
1–2	172 (47.6)	71 (31.4)	101 (74.8)	
3–4	189 (52.4)	155 (68.6)	34 (25.2)	
Hypertension, n (%)	207 (57.3)	130 (57.5)	77 (57)	0.928
Diabetes, n (%)	55 (15.2)	33 (14.6)	22 (16.3)	0.665
Presence of gross hematuria, n (%)	251 (69.5)	164 (72.6)	87 (64.4)	0.105
History of bladder cancer, n (%)	66 (18.3)	53 (23.5)	13 (9.6)	0.001
Baseline eGFR (ml/min/1.73 m ²)	72.3 (60.6, 85.5)	69.7 (60.1, 83.2)	75.7 (63, 88.3)	0.007
Baseline eGFR category ^b , n (%)				0.040
1 (eGFR ≥90)	60 (16.6)	31 (13.7)	29 (21.5)	
2 (60 ≤ eGFR < 90)	220 (61)	139 (61.5)	81 (60)	
3a (eGFR <60)	81 (22.4)	56 (24.8)	25 (18.5)	
Tumor laterality, n (%)				0.674
Right	187 (51.8)	119 (52.7)	68 (50.4)	
Left	174 (48.2)	107 (47.3)	67 (49.6)	
Tumor size (cm)	3.2 (2.2, 4.5)	3.0 (2.0, 4.1)	4.0 (2.5, 5.1)	<0.001
Tumor site, n (%)				<0.001
Renal calyx/pelvis	196 (54.3)	141 (62.4)	55 (40.7)	
Ureter	133 (36.8)	66 (29.2)	67 (49.6)	
Both	32 (8.9)	19 (8.4)	13 (9.6)	
Multifocal tumor, n (%)	59 (16.3)	40 (17.7)	19 (14.1)	0.367
Hydroureteronephrosis, n (%)	162 (44.9)	87 (38.5)	75 (55.6)	0.002
cT stage, n (%)				<0.001
T _{a-1}	239 (66.2)	192 (85)	47 (34.8)	
T ₂₋₄	122 (33.8)	34 (15)	88 (65.2)	
cN stage, n (%)				0.399
N ₀	344 (95.3)	217 (96)	127 (94.1)	
N ₁₋₂	17 (4.7)	9 (4)	8 (5.9)	
Robotic approach, n (%)	278 (77)	186 (82.3)	92 (68.1)	0.002
Bladder cuff management, n (%)				0.869
Standard excision	271 (75.1)	169 (74.8)	102 (75.6)	
Other technique ^c	90 (24.9)	57 (25.2)	33 (24.4)	
Tumor grade ^d , n (%)				0.159
Grade 1	56 (15.5)	40 (17.7)	16 (11.9)	
Grade 2	24 (6.7)	2 (0.9)	22 (16.3)	
Grade 3	270 (74.8)	176 (77.9)	94 (69.6)	
Unknown	11 (3.0)	8 (3.5)	3 (2.2)	
Follow-up duration (mo)	14 (7, 27)	14.3 (4, 29)	14 (9.8, 26)	0.352
Postoperative eGFR (ml/min/1.73 m ²)				
POD 1 (n = 310)	53.3 (44.8, 65.5)	49.6 (42.4, 56.9)	60.6 (49.2, 76.1)	<0.001
At discharge (n = 337)	53.2 (43.5, 66)	51.1 (41.9, 59.4)	60.2 (47.6, 74.8)	<0.001
3 mo (n = 361)	52.0 (42.5, 64.9)	50.5 (41.9, 60.4)	57.8 (45.2, 70.2)	<0.001
3 mo eGFR <50, n (%)	158 (43.8)	112 (49.6)	46 (34.1)	0.004
12 mo (n = 160)	49.9 (41.5, 60.9)	46.9 (39.2, 56)	53.9 (45.1, 65.8)	0.001
Last follow-up (n = 306)	52.6 (43.2, 65.9)	48.5 (40, 59)	58.7 (47.4, 73.7)	<0.001
Tumor recurrence, n (%)	107 (29.6)	78 (34.5)	29 (21.5)	0.009
Time to recurrence (mo)	6 (3, 10.3)	6 (3, 9.8)	7 (6, 12)	0.135
Tumor metastasis, n (%)	41 (11.4)	32 (14.2)	9 (6.7)	0.062
Time to metastasis (mo)	7 (4.5, 16)	6.8 (4, 12.7)	15 (6.8, 22)	0.123
Cancer-specific mortality, n (%)	25 (6.9)	22 (9.7)	3 (2.2)	0.007
Overall mortality, n (%)	36 (10)	31 (13.7)	5 (3.7)	0.002

ASA = American Society of Anesthesiologists; BMI = body mass index; eGFR = estimated Glomerular Filtration Rate; POD = postoperative day; RNU = radical nephroureterectomy.

Medians (interquartile range) or frequencies (proportions) are displayed for continuous and categorical variables, respectively.

^a p values compare the training set with the validation set using Mann-Whitney U test, chi-square test, or exact Fisher test, depending on whether the variable is continuous or categorical.

^b According to the Kidney Disease: Improving Global Outcomes guidelines for chronic kidney disease.

^c Pluck technique (n = 0), stripping (n = 2), transurethral resection (n = 24), intussusception (n = 2), Hem-o-Lok closure (n = 36), EndoGIA/Ligasure (n = 1), not removed because of contemporary cystectomy (n = 16), and unknown (n = 9).

^d According to the result of post-RNU pathological examination.

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(44.9%), 87 (38.5%), and 75 (55.6%) cases of the whole, training, and validation sets, respectively. The median (IQR) baseline eGFRs were, respectively, 72.3 (60.6, 85.5), 69.7 (60.1, 83.2), and 75.7 (63, 88.3) ml/min/1.73 m². Overall, 158 (43.8%) patients experienced eGFR reduction to <50 ml/min/1.73 m² at 3-mo postoperative follow-up, with 112 (49.6%) and 46 (34.1%) patients for the training and the validation group, respectively. The median follow-up for survivors was 14 (7, 27) mo. Overall, 110 patients were followed up for ≥24 mo. The overall mortality and cancer-specific mortality rates for the whole cohort were 10% and 6.9%, respectively.

3.2. Model development, performance evaluation, and external validation

Four preoperative factors (age, body mass index [BMI], preoperative eGFR, and hydroureteronephrosis) were screened by a LASSO regression as important features for scoring function development in the training set (Supplementary Fig. 1 and 2). An RCS analysis showed a nonlinear association between the variable of age and the outcome ($p < 0.05$), and the risk function demonstrated an inflection point at 70 yr, then age was transformed into a binary variable. A nonlinear association was not detected for the other two continuous factors (both $p > 0.05$; Supplementary Fig. 3).

A model that incorporated the above independent predictors was developed (see the Supplementary material for the formula of risk score calculation) and presented as the nomogram (Fig. 2) with an area under the curve (AUC) of 0.771 (95% CI: 0.711–0.831; Fig. 3A), which was confirmed to be 0.759 via bootstrapping validation. External validation of the nomogram revealed an AUC of 0.773 (95% CI: 0.697–0.850; Fig. 3B).

The calibration plot showed that the predicted probability of postoperative eGFR <50 ml/min/1.73 m² had concordance to that of the observed frequency, with most predictions within a mean absolute error of 0.031 in the training set (Fig. 3C) and 0.072 in the external validation set (Fig. 3D).

The Hosmer-Lemeshow goodness of fit test revealed $p = 0.16$ (the training set) and 0.263 (the external validation set), indicating nonsignificant miscalibration.

3.3. Clinical usefulness

The DCA for the nomogram is presented in Figure 4A. The DCA indicated that when the threshold probability ranges from 0 to 0.83, the nomogram adds more net benefit than the “treat all” or “treat none” strategies, which means that our nomogram could bring benefits to patients in practice.

Stratification of the high-risk probability for 1000 samples was predicted on the clinical impact curve (Fig. 4B). The predictive high-risk number was close to the actual number of positive cases when the threshold probability was >0.3. At this time, the cost-to-benefit ratio was 2:5. An optimal risk score cutoff value of 0.583 was achieved based on the principle of achieving a maximum Youden index in the training set. The relevant specificity was 0.81 and sensitivity 0.60. The patients were classified into low- and high-risk groups according to the optimal cutoff value. The high-risk group had a greater likelihood of postoperative eGFR <50 ml/min/1.73 m² (risk ratio [RR]: 2.618, 95% CI: 1.805–3.798 in the training set and RR: 1.604, 95% CI: 1.035–2.487 in the validation set). Notably, the difference in both overall death risk (HR: 4.59, 95% CI: 2.28–9.23, $p < 0.001$; Fig. 5A) and cancer-specific death risk (HR: 5.19, 95% CI: 2.24–12.04, $p < 0.001$; Fig. 5B) was statistically significant between the two groups for the combined training and validation set. Besides, the significance of difference in death risk between the two groups remained in the subgroup of pT_{2–4}N_x or TanyN_{1–2} patients (overall death risk, HR: 4.58, 95% CI: 2.17–9.65, $p < 0.001$; cancer-specific death risk, HR: 5.79, 95% CI: 2.36–14.25, $p < 0.001$; Fig. 5C and 5 D).

4. Discussion

In this retrospective global, multicenter, development and validation study, we constructed a clinical signature-based

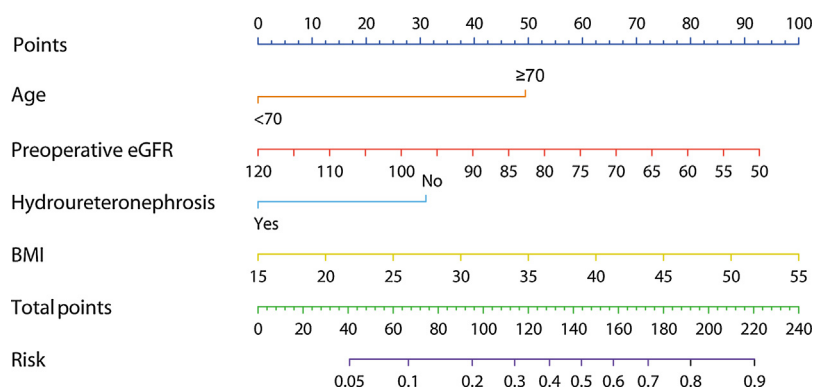


Fig. 2 – Nomogram for the prediction of postoperative eGFR <50 ml/min/1.73 m² after RNU based on a multivariable model. Instructions for utilization: locate the patient's preoperative eGFR on the corresponding axis. Draw a line straight downward to the point axis to determine how many points toward the probability of postoperative eGFR <50 ml/min/1.73 m² the patient receives for the preoperative eGFR. Repeat the process for each additional predictor. Add the points for each of the variables. Locate the final sum on the total points axis. Draw a line straight up to find the patient's risk probability. BMI = body mass index; eGFR = estimated glomerular filtration rate; RNU = radical nephroureterectomy.

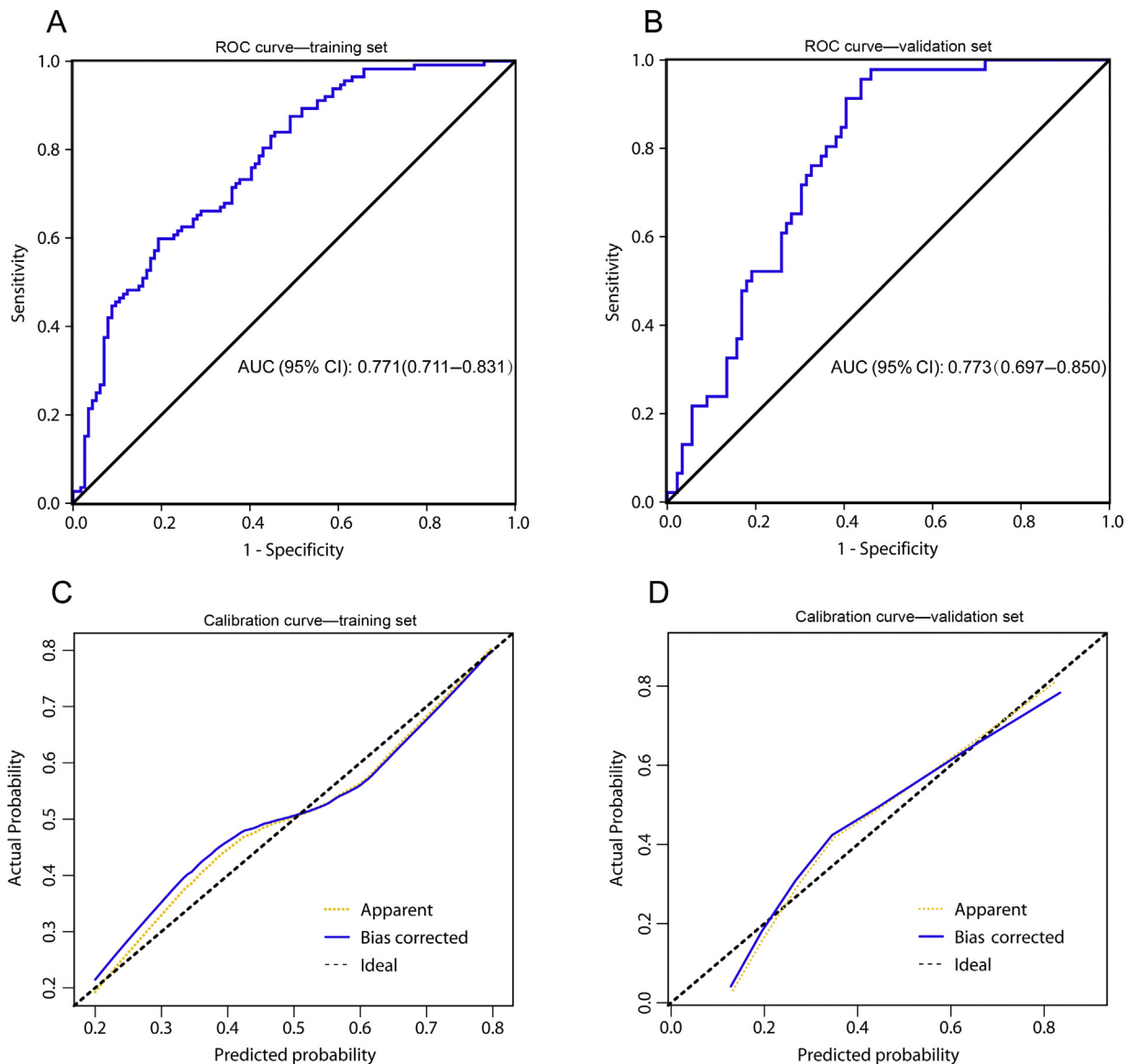


Fig. 3 – ROC curve and calibration plot of the model. (A and B) ROC curves predicted probabilities from the nomogram in the training set and the external validation set. (C and D) Calibration curves for the nomogram to predict probability of postoperative eGFR <50 ml/min/1.73 m² for the training set and the external validation set. AUC = area under the curve; CI = confidence interval; eGFR = estimated glomerular filtration rate; ROC = receiver operating characteristics.

nomogram for the preoperative individualized prediction of renal function insufficiency for cisplatin-based adjuvant chemotherapy in UTUC patients after RNU. The model successfully stratified patients according to their risk of renal function insufficiency for cisplatin use as well as mortality. Incorporating clinical risk factors into an easy-to-use nomogram facilitates the preoperative individualized risk prediction and optimum timing of perioperative cisplatin-based chemotherapy.

The nomogram AUC of 0.771 (95% CI: 0.711–0.831) in the development cohort confirmed good discrimination ability. However, application of a prediction model is always limited by the type of validation or the lack of external validation [15,16]. In this study, to improve the generalizability of

the model's performance, we constructed a geographically independent validation set comprising 135 patients from four different institutions in four different non-USA countries. The difference in the distribution of important variables (demographics, predictors, and outcome; Table 1) reveals that the validation set comprises populations that are “plausibly related” to the development cohort. The achieved predictive accuracy of the nomogram in the external validation cohort was comparable with a slight improvement (AUC of 0.773, 95% CI: 0.697–0.850), although the outcome positivity was much lower in the validation cohort (34.1% vs 49.6%), which suggested a robust prediction model sufficiently capturing informative predictive factors. The calibration plot showed good agreement

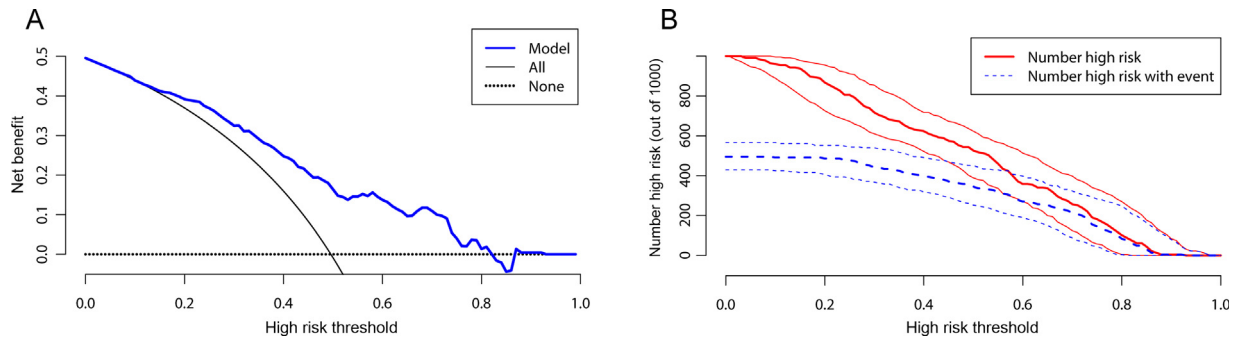


Fig. 4 – Decision curve and clinical impact curve for the predictive nomogram. (A) The net benefits were measured at different threshold probabilities. The blue line represents the predictive nomogram. The solid line represents the assumption that all patients have the study outcomes and no intervention is done. The dotted line represents the assumption that no patients have the study outcomes and no intervention is done. **(B)** Clinical impact curve to predict the high-risk number for a population size of 1000. The red curve shows the predicted high-risk number at different threshold probabilities and the blue curve represents the actual high-risk patients.

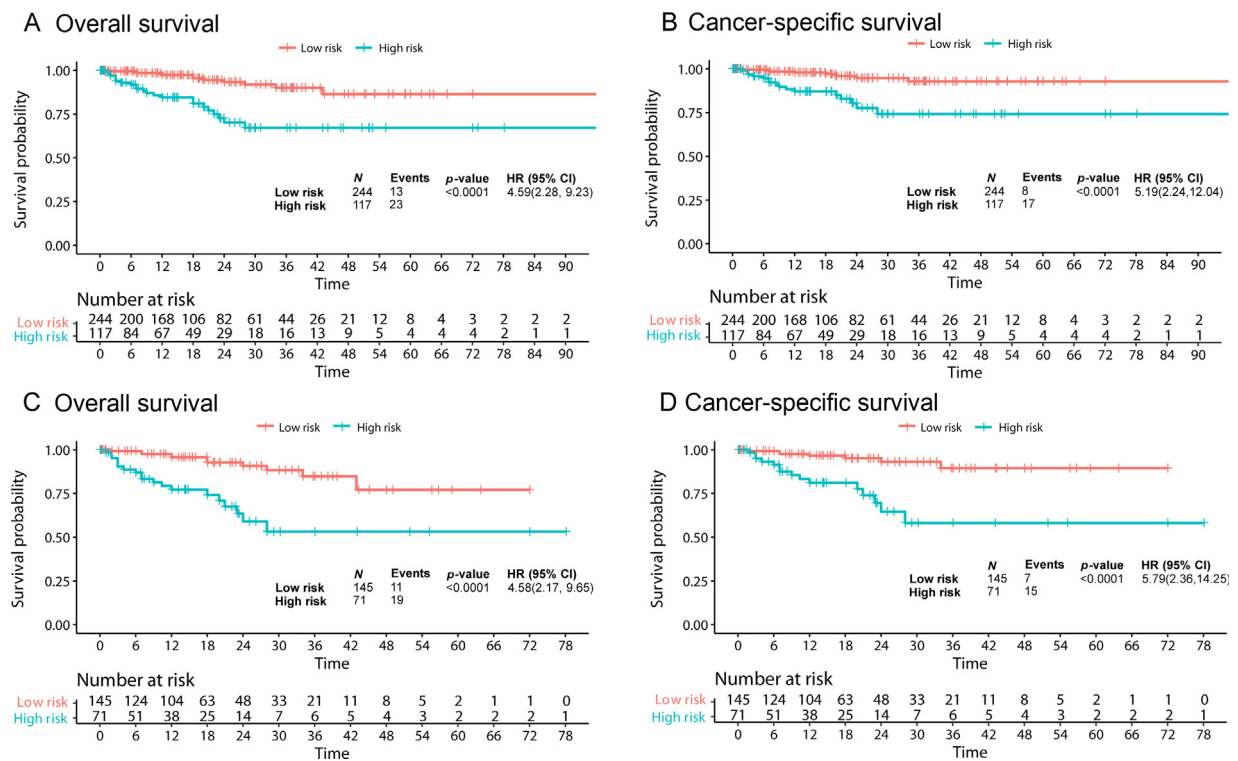


Fig. 5 – Kaplan-Meier estimates with Cox regression analysis for overall and cancer-specific survival in RNU patients with low and high risk for postoperative eGFR <50 ml/min/1.73 m². (A and B) The whole group combined the training set and validation set. **(C and D)** The subgroup of pT2-4Nx or TanyN1-2 patients. CI = confidence interval; eGFR = estimated glomerular filtration rate; HR = hazard ratio; RNU = radical nephroureterectomy.

between predicted probability and observed frequency. Last but not least, robustness of the model was confirmed via a sensitivity analysis.

Many clinical factors are found to be associated with worse renal functional outcomes following RNU [6,17,18], but there is a high risk of overfitting, multicollinearity, and interaction in predictive modeling of these data, characterized by reduced significance of the predictor when applied to an independent dataset [19]. The LASSO penalization for optimal variable selection not only surpasses the method of selecting predictors on the basis of the strength of their univariable association with outcome, but also enables the

panel of candidate features to be converted into a combined signature [20,21]. In the present study, we have applied this method successfully and identified four clinical items: age at surgery, BMI, preoperative eGFR, and hydroureteronephrosis of the operated kidney. Additionally, linearity may not always hold between the relationship of a continuous independent variable and the dependent variable. To address this issue, an RCS analysis was adopted [22], which has recently been used to study the relationship between survival/outcome and treatment in patients with cancer [23,24]. In this study, it uncovered nonlinearity of the relationship between patient age and the predicted

outcome with an inflection point at 70 yr (if incorporated as a continuous variable, the AUC decreased to 0.736; data not shown), a threshold in concordance to others despite their relatively arbitrary classification [5,6,17].

Contrary to what is reported in previous studies [6,17,25], our results demonstrated that BMI was an independent risk factor for renal function decline after RNU. Although the relationship between obesity and CKD progression is complex, its negative effect has been revealed from large prospective series [26,27]. Not surprisingly, preoperative renal function status is predominant for postoperative functional preservation. Actually, split renal function of the operated kidney by nuclear renal scans could be more accurate to reflect the function loss, but it is not routinely obtained before the surgery of RNU. Nevertheless, some alternative factors related to the split renal function such as tumor size, multifocality, ureteral location, and hydronephrosis are more practical to use in our clinical practice. Owing to the overlapping effect between these variables, the results derived from common multivariable regression analysis in previous studies are not consistent [6,17,18,25]. In our study, only the hydroureteronephrosis was demonstrated to be a predictive factor after adjusting for others with improved methodology. Finally, the model shows additional value in prognostic evaluation. A Cox regression analysis indicated that the risk cutoff score based on our model was a significant predictive factor distinct from the pathological traits. In this regard, should these patients be considered for neoadjuvant chemotherapy, it will be important to maximize preoperative renal function via ureteral stents or percutaneous drainage when indicated.

The clinical usefulness of the nomogram was also evaluated to facilitate decision-making on further intervention. When categorized patients into low- and high-risk groups on the basis of the cutoff values derived from the nomogram; the high-risk group had a significantly greater probability of having a predicted outcome in both the training and the independent validation group. Furthermore, the present model showed better net benefit at a threshold probability of $\leq 83\%$. At this time, we still lack confidence to avoid overtreatment versus risk of disease progression, and there is recommendation to provide neoadjuvant cisplatin-based multidrug regimens, given the high likelihood (43.8%, as shown in Table 1) of being unable to offer cisplatin-based therapy in adjuvant setting, similar to previous reports [5,7,25,28]. Therefore, the prediction tool may be especially helpful for treatment selection in favorable surgical candidates classified at a lower risk for cisplatin ineligibility to skip neoadjuvant chemotherapy, and, on the contrary hand, for treatment selection of those who are at a high risk of renal insufficiency postoperatively and would probably benefit from neoadjuvant chemotherapy.

The main limitation of the present study lies in its retrospective nature despite the data collected in a structured form. It is still subject to selection and recall biases, and some unmeasured data (ie, split renal function) that were not retrospectively retrievable. Missing data on predictors led to cases being excluded and to a further

decrease in the power of our study. Moreover, we have to acknowledge that postoperative renal function is a time-sensitive outcome and eGFR fluctuations are common in the immediate postoperative period (within 1 mo). Nevertheless, the proper eGFR (median 3 mo, IQR 1.8–3 mo) was selected in this study to demonstrate cisplatin eligibility in the adjuvant setting if the patient indeed required systemic therapy. Finally, the follow-up duration was relatively short for survival analysis. Prospective randomized clinical trials with a large sample size are needed to acquire high-level evidence for clinical application in the future.

5. Conclusions

A nomogram incorporating immediately available clinical variables can accurately predict renal insufficiency for cisplatin-based adjuvant chemotherapy after minimally invasive RNU and may serve as a tool optimizing patient selection.

Author contributions: Linhui Wang had full access to all the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis.

Study concept and design: Wu, Autorino.

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Analysis and interpretation of data: Wu, Autorino.

Drafting of the manuscript: Wu.

Critical revision of the manuscript for important intellectual content: Autorino, Vecchia, Djaladat, Mari, Margulis.

Statistical analysis: Wu, Chen.

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Appendix A. Supplementary data

Supplementary material related to this article can be found, in the online version, at doi:<https://doi.org/10.1016/j.euf.2021.01.014>.

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